

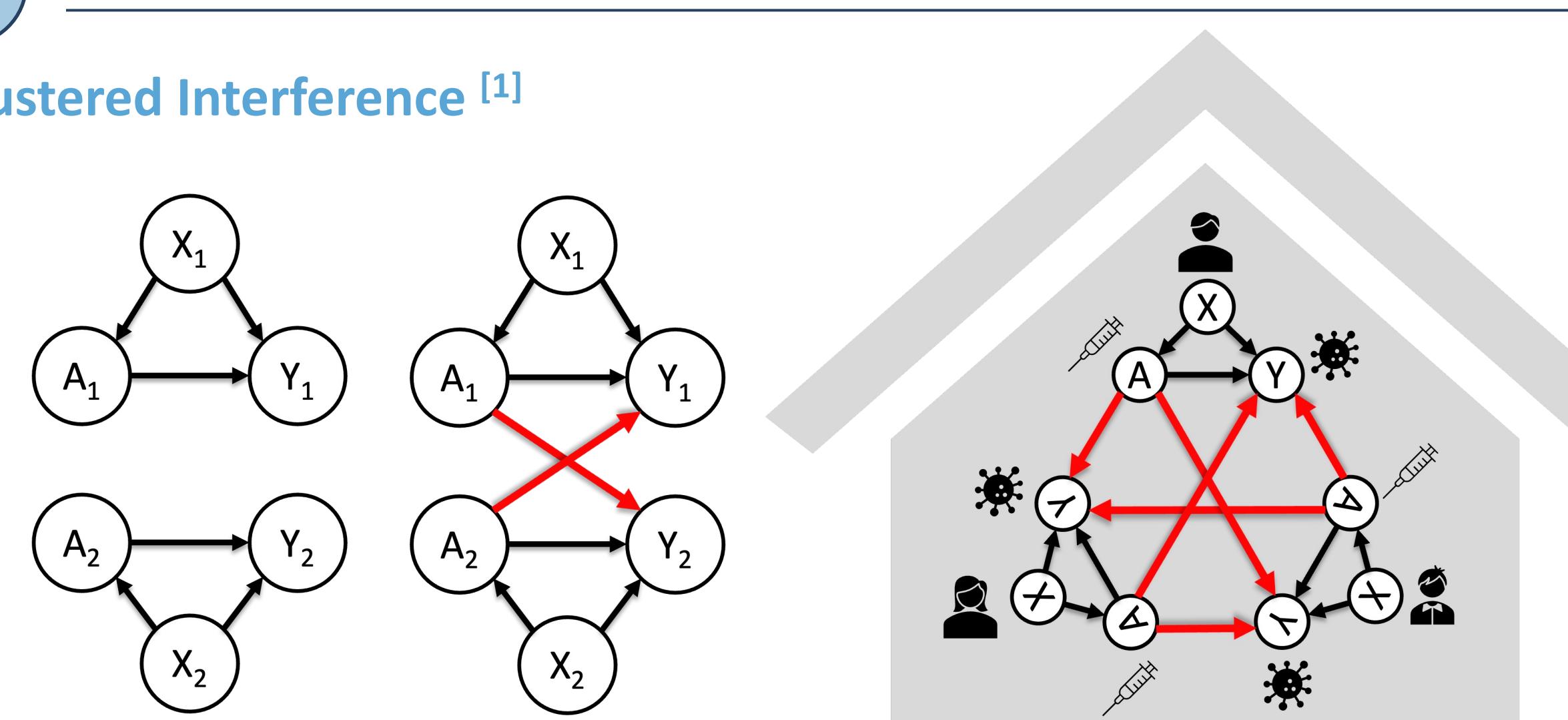


Chanhwa Lee¹, Donglin Zeng², Michael G. Hudgens¹

¹Department of Biostatistics, University of North Carolina at Chapel Hill, ²Department of Biostatistics, University of Michigan, Ann Arbor

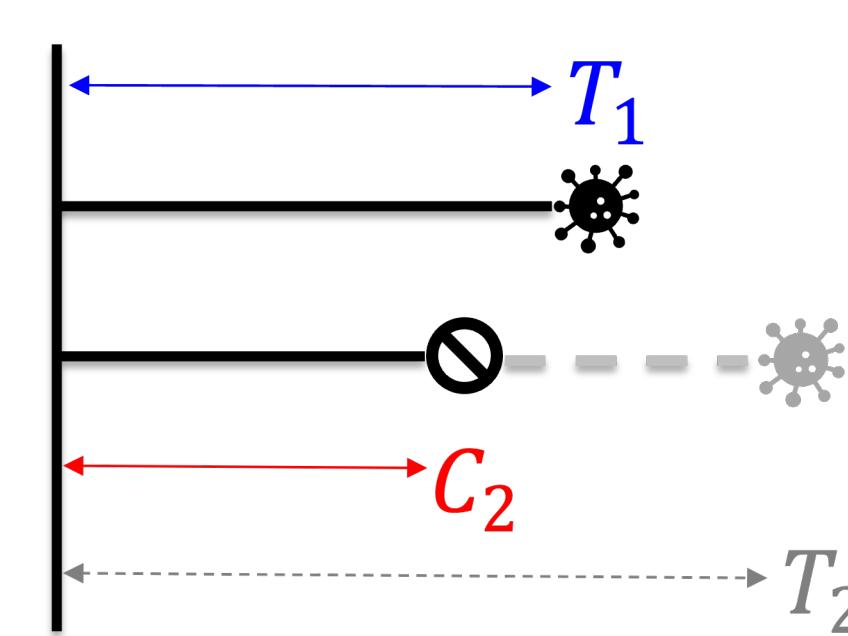
1 MOTIVATION

Clustered Interference [1]

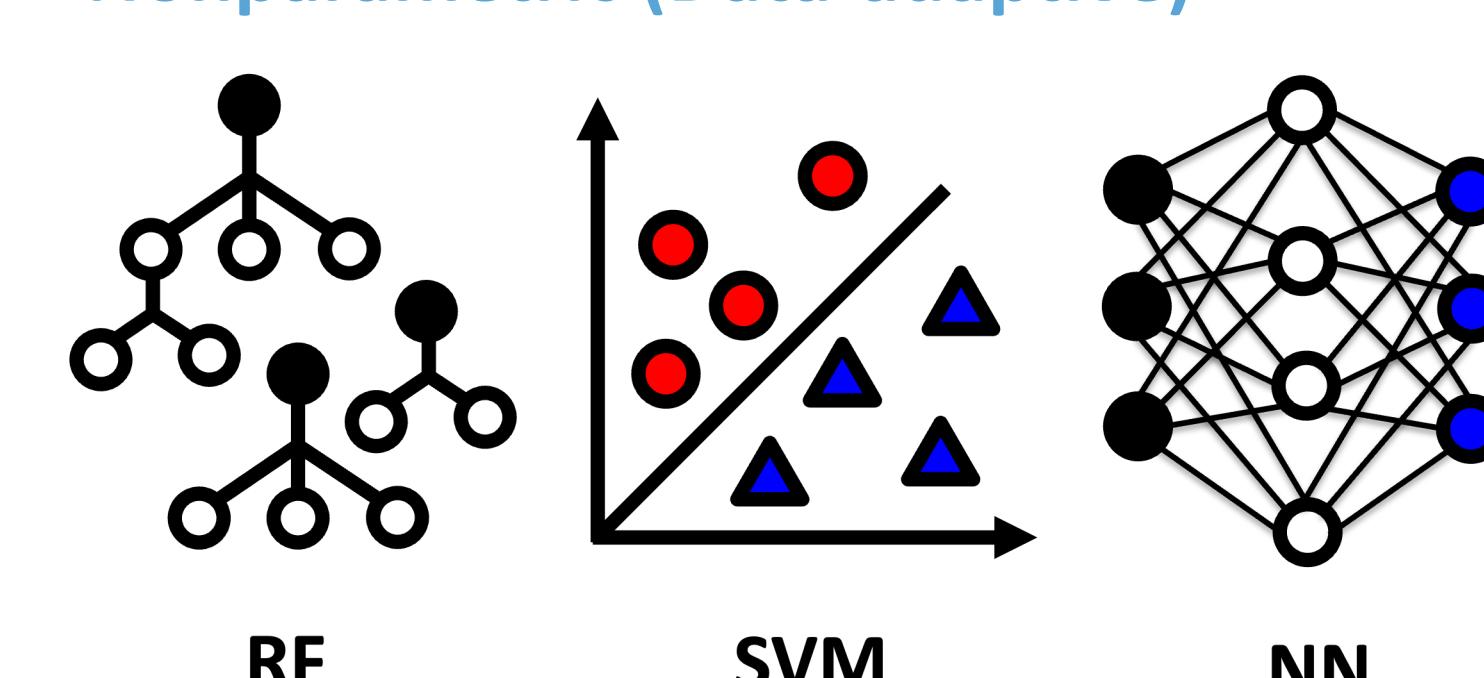


No interference

Right Censoring [2]



Nonparametric (Data-adaptive) [3]



2 BACKGROUND

Observed data

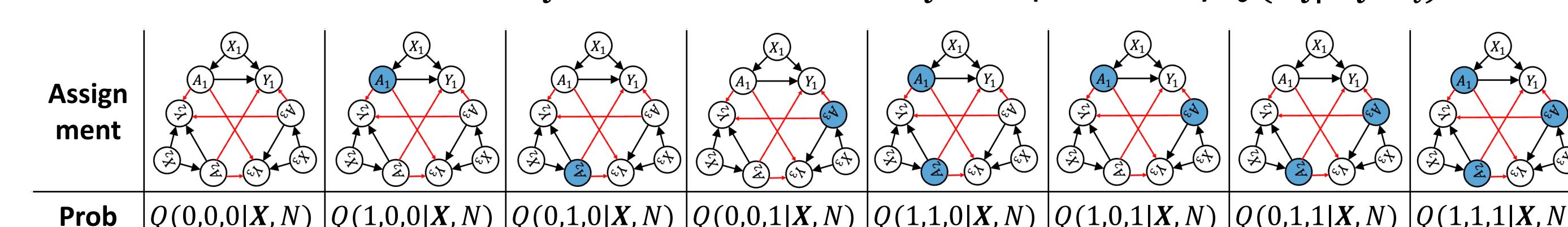
- Cluster $i \in \{1, \dots, m\}$, Unit $j \in \{1, \dots, N_i\}$
- $T_{ij} \in \mathbb{R}^+$: event time, $A_{ij} \in \{0,1\}$: treatment, $X_{ij} \in \mathbb{R}^p$: confounders
- $C_{ij} \in \mathbb{R}^+$: censoring time, $Y_{ij} = \min(T_{ij}, C_{ij})$, $\Delta_{ij} = 1(T_{ij} \leq C_{ij})$

Potential outcome

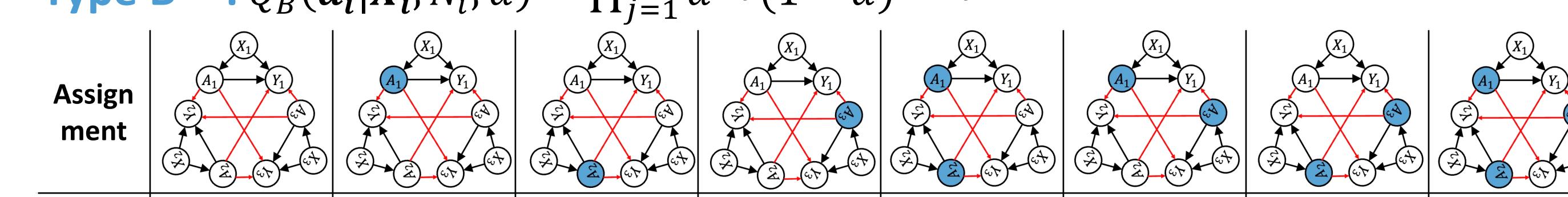
- $T_{ij}(\mathbf{a}_i)$: Potential event time for unit j in cluster i when the cluster i receives \mathbf{a}_i
- $T_{ij}(\mathbf{a}_i) = T_{ij}(a_{ij}, \mathbf{a}_{i(-j)})$, $\mathbf{a}_{i(-j)} = (a_{i1}, \dots, a_{i(j-1)}, a_{i(j+1)}, \dots, a_{iN_i})$
- No interference: $T_{ij}(a_{ij}, \mathbf{a}_{i(-j)}) = T_{ij}(a_{ij}, \mathbf{a}'_{i(-j)})$

3 STOCHASTIC POLICY

Definition $Q(\cdot | \mathbf{X}_i, N_i)$: probability distribution on $\{0,1\}^{N_i}$ such that cluster of size N_i with cluster-level covariate \mathbf{X}_i receives treatment \mathbf{a}_i with probability $Q(\mathbf{a}_i | \mathbf{X}_i, N_i)$



Type B [4]: $Q_B(\mathbf{a}_i | \mathbf{X}_i, N_i; \alpha) = \prod_{j=1}^{N_i} \alpha^{a_{ij}} (1-\alpha)^{1-a_{ij}}$



CIPS [5]: $Q_{CIPS}(\mathbf{a}_i | \mathbf{X}_i, N_i; \delta) = \prod_{j=1}^{N_i} \pi_{ij,\delta}^{a_{ij}} (1 - \pi_{ij,\delta})^{1-a_{ij}}$

- Propensity score of unit j in cluster i : $\pi_{ij} = P(A_{ij} = 1 | \mathbf{X}_i, N_i)$
- Shifted (counterfactual) propensity score: $\pi_{ij,\delta}$ from $\frac{\pi_{ij,\delta}}{1-\pi_{ij,\delta}} = \delta \times \frac{\pi_{ij}}{1-\pi_{ij}}$

Literature cited

- Hudgens, M. G. and Halloran, M. E. (2008). Toward causal inference with interference. *Journal of the American Statistical Association* 103, 832–842.
- Chakladar, S., Rosin, S., Hudgens, M. G., Halloran, M. E., Clemens, J. D., Ali, M., and Emch, M. E. (2022). Inverse probability weighted estimators of vaccine effects accommodating partial interference and censoring. *Biometrics* 78, 777–788.
- Park, C. and Kang, H. (2022). Efficient semiparametric estimation of network treatment effects under partial interference. *Biometrika* 109, 1015–1031.
- Tchetgen Tchetgen, E. J. and VanderWeele, T. J. (2012). On causal inference in the presence of interference. *Statistical Methods in Medical Research* 21, 55–75.
- Lee, C., Zeng, D., and Hudgens, M. G. (2024). Efficient Nonparametric Estimation of Stochastic Policy Effects with Clustered Interference. *Journal of the American Statistical Association*
- Chernozhukov, V., Chetverikov, D., Demirer, M., Duflo, E., Hansen, C., Newey, W., and Robins, J. (2018). Double/debiased machine learning for treatment and structural parameters. *The Econometrics Journal* 21, C1–C68.

4 ESTIMANDS

• Expected Overall Risk by time τ under policy Q

$$\mu(\tau; Q) = \mathbb{E} \left\{ \frac{1}{N_i} \sum_{j=1}^{N_i} \sum_{\mathbf{a}_i \in \mathcal{A}(N_i)} \mathbb{1}(T_{ij}(\mathbf{a}_i) \leq \tau) Q(\mathbf{a}_i | \mathbf{X}_i, N_i) \right\}$$

• Expected Risk by time τ when treated under policy Q

$$\mu_1(\tau; Q) = \mathbb{E} \left\{ \frac{1}{N_i} \sum_{j=1}^{N_i} \sum_{\mathbf{a}_{i(-j)} \in \mathcal{A}(N_i-1)} \mathbb{1}(T_{ij}(1, \mathbf{a}_{i(-j)}) \leq \tau) Q(\mathbf{a}_{i(-j)} | \mathbf{X}_i, N_i) \right\}$$

• $OE(\tau; Q, Q') = \mu(\tau; Q) - \mu(\tau; Q')$: compares two policies overall

E.g. Difference in overall COVID19 risks when 50% vs. 30% of neighbors vaccinated

• $DE(\tau; Q) = \mu_1(\tau; Q) - \mu_0(\tau; Q)$: effect of treatment under policy Q

E.g. Vaccine effect on COVID19 when 50% of neighbors vaccinated

• $SE_0(\tau; Q, Q') = \mu_0(\tau; Q) - \mu_0(\tau; Q')$: compares risk when untreated

E.g. Unvaccinated unit's COVID19 risks when 50% vs. 30% of neighbors vaccinated

5 METHOD

1. Full data estimation equation from nonparametric EIF

$$\varphi_{ij}^{F_*}(\tau; \mathbf{Z}_i) := \sum_{\mathbf{a}_i \in \mathcal{A}(N_i)} \{w_j(\mathbf{a}_i, \mathbf{X}_i, N_i) + \phi_j(\mathbf{A}_i, \mathbf{X}_i, N_i; \mathbf{a}_i)\} \mathbb{P}(T_{ij} \leq \tau | \mathbf{A}_i = \mathbf{a}_i, \mathbf{X}_i, N_i) \\ + \mathbb{P}(\mathbf{A}_i | \mathbf{X}_i, N_i)^{-1} w_j(\mathbf{A}_i, \mathbf{X}_i, N_i) \{\mathbb{1}(T_{ij} \leq \tau) - \mathbb{P}(T_{ij} \leq \tau | \mathbf{A}_i, \mathbf{X}_i, N_i)\} - \Psi(\tau; \mathbf{w})$$

2. Augmented-IPCW estimating equation

$$0 = \frac{1}{m} \sum_{i=1}^m \frac{1}{N_i} \sum_{j=1}^{N_i} \left[\frac{\Delta_{ij}}{S_{ij}^C(Y_{ij} | \mathbf{A}_i, \mathbf{X}_i, N_i)} \varphi_{ij}^{F_*}(\tau; \mathbf{Z}_i) + \int_0^\infty \frac{\mathbb{E}\{\varphi_{ij}^{F_*}(\tau; \mathbf{Z}_i) | T_{ij} \geq r, \mathbf{A}_i, \mathbf{X}_i, N_i\}}{S_{ij}^C(r | \mathbf{A}_i, \mathbf{X}_i, N_i)} dM_{ij}^C(r) \right]$$

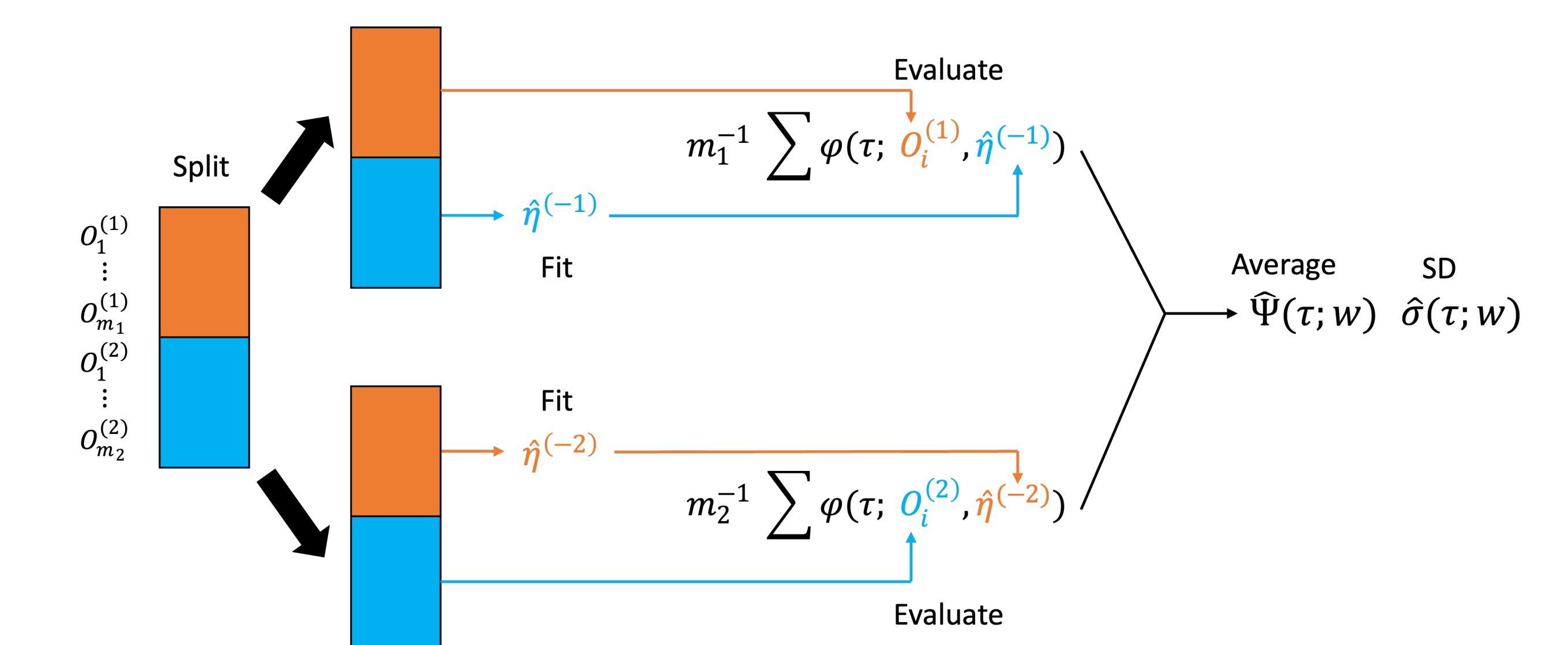
$$\varphi_{ij}(\tau; \mathbf{O}_i) = \sum_{\mathbf{a}_i \in \mathcal{A}(N_i)} \text{OR}_{ij}(\tau; \mathbf{O}_i, \mathbf{a}_i) + \text{IPCW-BC}_{ij}(\tau; \mathbf{O}_i) + \text{AUG}_{ij}(\tau; \mathbf{O}_i)$$

$$\text{OR}_{ij}(\tau; \mathbf{O}_i, \mathbf{a}_i) = \{w_j(\mathbf{a}_i, \mathbf{X}_i, N_i) + \phi_j(\mathbf{A}_i, \mathbf{X}_i, N_i; \mathbf{a}_i)\} F_{ij}^T(\tau | \mathbf{A}_i = \mathbf{a}_i, \mathbf{X}_i, N_i),$$

$$\text{IPCW-BC}_{ij}(\tau; \mathbf{O}_i) = \frac{w_j(\mathbf{A}_i, \mathbf{X}_i, N_i)}{\mathbb{P}(\mathbf{A}_i | \mathbf{X}_i, N_i)} \left\{ \frac{\Delta_{ij}}{S_{ij}^C(Y_{ij} | \mathbf{A}_i, \mathbf{X}_i, N_i)} \mathbb{1}(Y_{ij} \leq \tau) - F_{ij}^T(\tau | \mathbf{A}_i, \mathbf{X}_i, N_i) \right\},$$

$$\text{AUG}_{ij}(\tau; \mathbf{O}_i) = \frac{w_j(\mathbf{A}_i, \mathbf{X}_i, N_i)}{\mathbb{P}(\mathbf{A}_i | \mathbf{X}_i, N_i)} \int_0^\tau \frac{S_{ij}^T(r | \mathbf{A}_i, \mathbf{X}_i, N_i) - S_{ij}^T(\tau | \mathbf{A}_i, \mathbf{X}_i, N_i)}{S_{ij}^C(r | \mathbf{A}_i, \mathbf{X}_i, N_i) S_{ij}^T(r | \mathbf{A}_i, \mathbf{X}_i, N_i)} dM_{ij}^C(r),$$

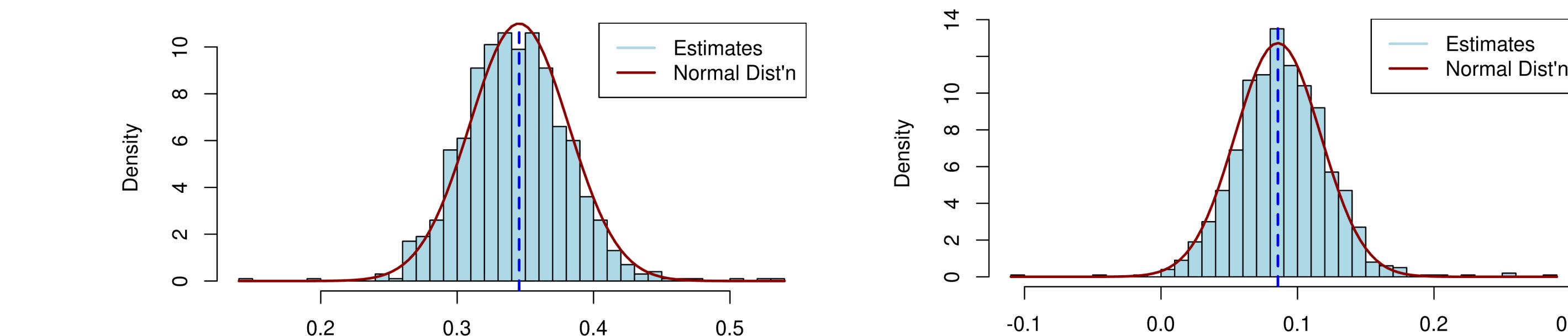
3. Sample Splitting [6] & Nonparametric Nuisance Function Estimation



6 RESULTS

Theory

- Consistent & Asymptotically Normal & Weak Convergence to Gaussian Process



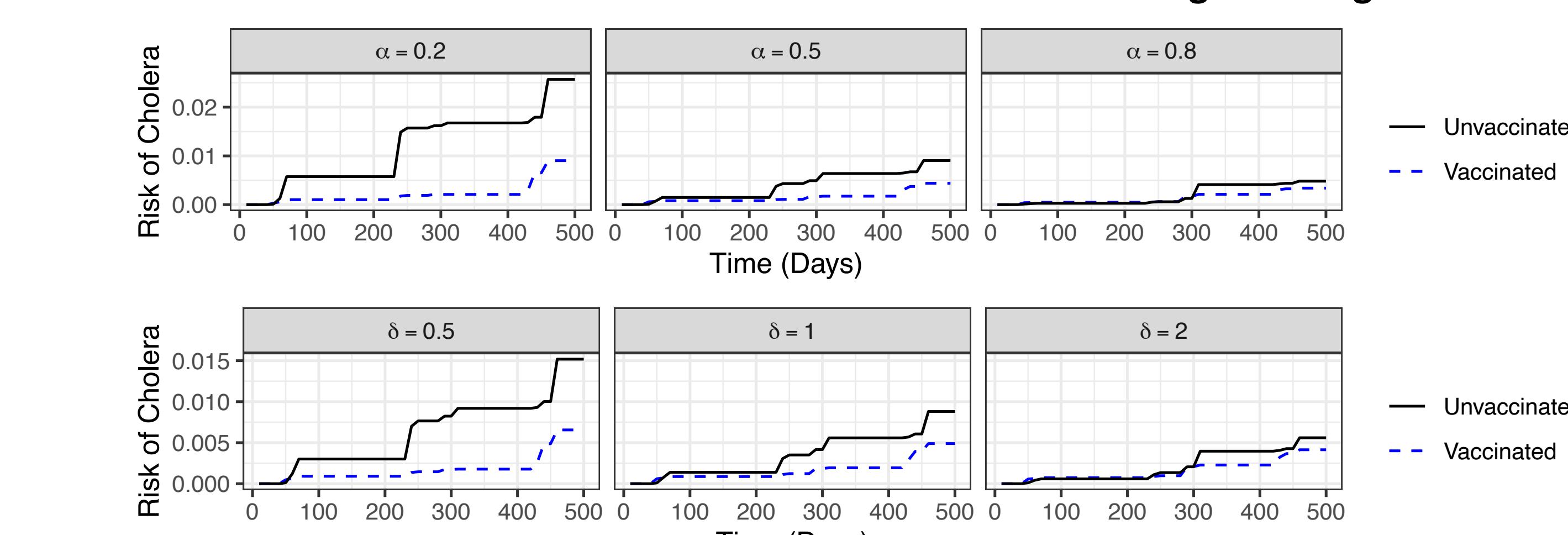
Simulation

- Flexible data-adaptive nuisance function estimation

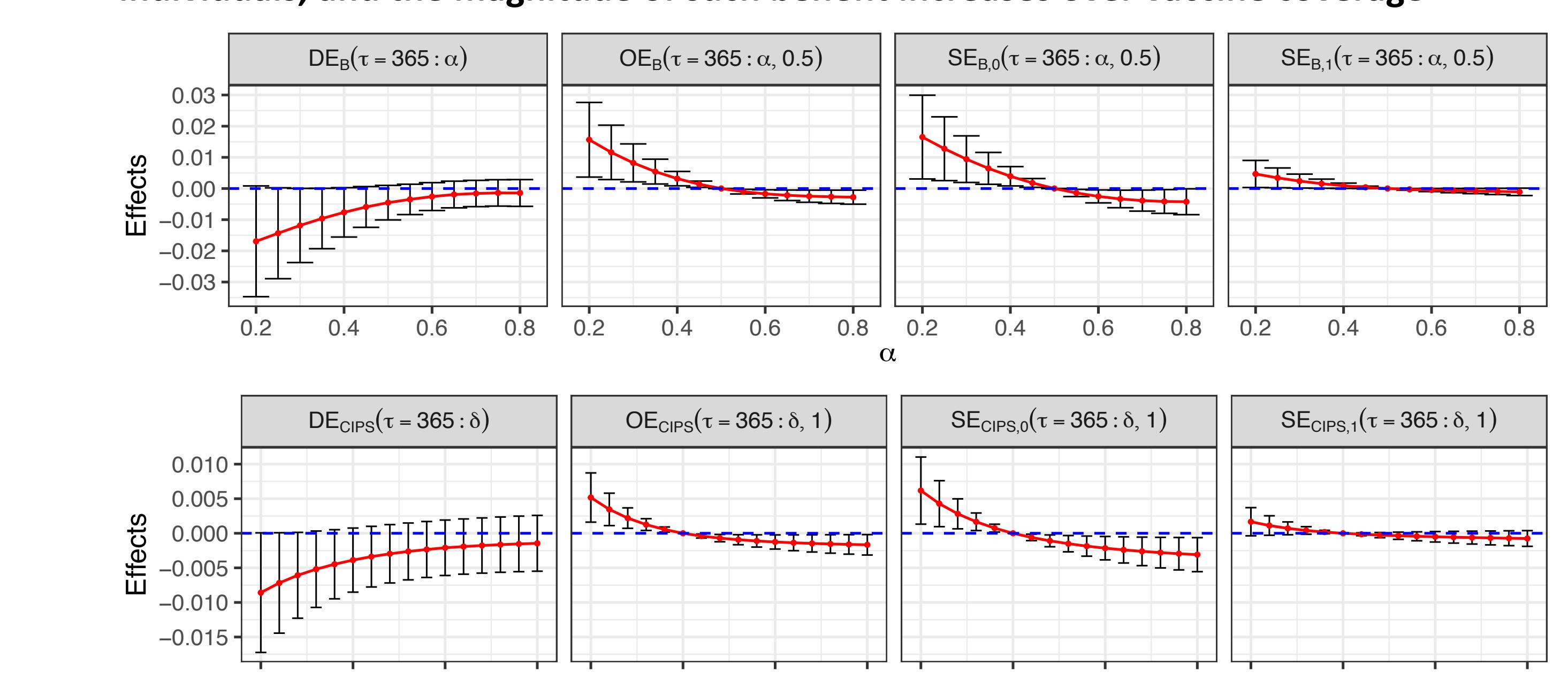
Estimand	Truth	Nonparametric (SL, RSF)			Parametric (GLM, Cox)			RMSE Ratio			
		Bias	RMSE	ASE	ESE	Cov %	Bias	RMSE	ASE	ESE	Cov %
$\mu_{CIPS}(0.3; 0.5)$	0.345	-0.003	0.001	0.036	0.038	94.1%	-0.024	0.002	0.038	0.041	87.4%
$\mu_{CIPS,1}(0.3; 0.5)$	0.151	-0.001	0.001	0.025	0.027	93.9%	-0.016	0.001	0.027	0.028	88.3%
$\mu_{CIPS,0}(0.3; 0.5)$	0.524	-0.004	0.004	0.058	0.063	94.3%	-0.032	0.005	0.063	0.066	90.4%
$DE_{CIPS}(0.3; 0.5)$	-0.373	0.002	0.004	0.058	0.064	94.1%	0.015	0.005	0.063	0.067	93.9%
$SE_{CIPS,1}(0.3; 0.5, 1)$	0.021	0.001	0.001	0.021	0.024	94.9%	-0.003	0.001	0.023	0.025	95.1%
$SE_{CIPS,0}(0.3; 0.5, 1)$	0.030	0.006	0.003	0.051	0.059	95.5%	-0.004	0.004	0.054	0.059	95.5%
$OE_{CIPS}(0.3; 0.5, 1)$	0.086	0.003	0.001	0.031	0.034	94.8%	-0.005	0.001	0.033	0.038	94.5%

Application (Cholera Vaccine Study)

- Beneficial direct effect of vaccination at lower vaccine coverage
- Beneficial indirect effect from vaccinated → unvaccinated at high coverage



- Unvaccinated individuals can benefit from spillover effects from vaccinated individuals, and the magnitude of such benefit increases over vaccine coverage



7 Discussion

- Inference about treatment effects under clustered interference and censoring
- Can be applied to any stochastic treatment allocation policy
- Data-adaptive estimation with robust correction to yield CAN estimator

Further information

Please address questions or comments to Chanhwa Lee at chanhwa@email.unc.edu.

Funding

This work was supported by a grant from the National Institutes of Health, USA (NIH grant NIH R01 AI085073).